## Consistent correction <br> of census and vital registration data for Thailand, 1960-80

Norman Y. Luther, Neramit Dhanasakdi, and Fred Arnold

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#### Abstract

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#### Abstract

A new procedure for the simultaneous and consistent corre. tion of two or more censuses and intercensal registered births and deaths is applied to Thailand's censuses of 1960, 1970, and 1980, and to registered births and deaths for the period 1960-80. The procedure begins with a set of preliminary correction factors that are inconsistent because they are derived from a sariety of existing sources. The procedure then identifies the set of consistent correction factors that are "closest" to the preliminary tacturs.

The results show marked underenumeration of the ()-4 age group in each consus, especiatly the 1906 census in which age rounding was apparently presalent. Age exaggeration at the oldest ages is also indicated, at least in the 1970 and 1480 censuses. Birth registration completemess rose during: 1960)-80, from about 82 to 85 percent for males and from about 78 to 45 percent for pemales. Nevertheless, death registration completeness for both sexes howered within the th to 6 percent range, dedining slightly until the carly 1970 , when it leveled off ar began torise. The age proup 0 - 4 had the greatest underregistration of deathe, and underregistration at this age apparently increased throughout the period. Athough life enpectancy incredsed for both seeses during the perind, temale superiarty in life evpectancy alsonemeded. The werage number of sars of life expectancy is estimated to hase increased from  for males in 1975 sio.


The Thailand censuse of 1960 , 1970 , and 1980 and vital registration during the period doth-so have been the subject of warious estimates of census undercount, difusted consus age-sex distributions, underregistration of births and deathe, intercensal life tables, and other demographic variables, many of which are mentioned helow. Howerer, these estimates, as a whole contain many incomaistenctes. The purpose of this paper is to provide a new, con sistent set of estimate for the ese data, making selective use of the presious entimater.

Io acomplish this purpuse, a new procedure for the simulaneous and consistent earrection of two or more remsunes and intereensal registered births and deathe (Luther and Rethertord 1986a, l9860) is applied to Thailand's cemsuses of 1900, 19F0, and 1980 and wo the registered births and deaths for the intervening period of 1 to (1) $8(0)$. Specificalls, for cach sex the procedure provides estimates of (1) correction factors for intercensal registered births, (2) age-specific correction iactors for intercensal registered deaths, (3) age-specific correction tactors for the age distribution at each census, and (4) life tables for the intercensal periods, such that the corrected registered births and deaths, the corrected census age distributions, and the estimated life tables are mutually consistent. Furthermore, the correctedbrthes satisty a specified sex ratio. Consiste acy here means that the usual intercensal demographic balancing equations relating census age distribu-
tions, births, and deaths by age are exactly satisfied. Net international migration was negligible for the period 1960-80, so that it may be ignored. The volume of fermanent immigration and emigration was small throughour this period. Temporary migration of contract workers to the Middle East began to increase in 1978, but it did not become substantial until after the 1980 census.

## METHOD

The procedure begins with a set of preliminary correction factors that are not necessarily consistent. These preliminary correction factors are derived from previous estimates and by existing methods of demographic analysis, as detailed below. Starting from this set of preliminary correction factors and the demographic data themselves, the new procedure specifies a "best," or optimal, set of final consistent correction factors that is "closest" to the set of preliminary correction tactors. The optimization procedure finds this closest set by use of the principice, from the mathematics of finitedimensional vector spaces, that there exists in a hyperplane a unique point of minimum distance from a fixed point not in the hyperphame. The methodology is explained in I.uther and Retherford (19i6a, 198(ob), We latter source gives details for the non-mathematically-indined reader. It is shown in those sources that, by atilizeg the concept of weighted distance, the procedure allows the analyst to weight the preliminary estimales aceording to the analyst's as essment of their dacuracy The use of "proportional weights," as described in the application of the procedure to South Korean data for 1970-80 (Luther and ketherford 1986a, 1986), will spred the correction evenly over the preliminary correction tactors. This is also appropriate for the Thatand application since we would assess all of the preliminary correction factors derived below to be of comparable dacuracy:

The three-census application to thailand given in this paper illustrates an important advantage of this procedure: that it may be used to correct consistently more than two consuses at a time, along with intercensa births and deaths. In contrast, most previous procedures correct just two censuses at a time (e.g., Demeny and Shorter 1968; Preston 1983), and that typically leads to inconsistent results for the intermediate censuses from the separate applications of the procedure. In particular, most previous analyses of Thailand data have involved just two successive censuses (e.g., Fulton 1979); Arnold and Phananiramai 1975; Pejaranonda, Arnold, and Hauser 1983).

Another advantage of the consistent correction procedure is that it allows the researcher great leewdy in the amount and type of structure that is imposed. For example, one may require the final age-specific correction factors to be the same from census to census, as with the Demeny-Shorter technique (Demeny and Shorter 1968) and the "variable $r$ " metliods (Preston
1983). No such assumptions are made for the Thai application given here, however. This latter, more flexible approach, which is often more realistic and appropriate, was adopted for the Korean application as well (Luther and Retherford 1986a, 1986b). Or one may compromise and assume that the final correction factors follow some trend. As an example of a different type of restriction that the procedure allows, we have fixed the sex ratio for the corrected Thai births. In fact, the only conditions that must be satisfied by the restrictions involsing the correction factors are that these restrictions be linear in the factors and hold when all factors are zero.

The procedure, however, has limitations. The primary one is that the final correction factess are sensitive to errors in the preliminary correction factors. Consequently, it is important to estimate the preliminary correction factors as accurately as prossible.

In any case, the final correction factors given by the procedure yield consistent adjustments of the data. The degree to which these adjustments are actually corrections depends on the atcuray of the preliminary correction factors.

## APIIICATION TO THALIAND, 1960-80

All censuses are treated as if the wocurred on 1 April, although the 1960 census was taten on 25 april. The four five year interensal periods that will be frequently mentioned are those separated be I April of the years 1960, 1965, 1970 . 1975 and 1980.

The input mto the procedure, and specificalls into the computer program (which may be whained trem the authers). consists of: (1) the 1960), 1970, and 1984) reperted census age distributions by sen and five-year age groups (open-ended at age 70 ), given in Tables 1 and 2 in the next section. with rach censun age distribution umiformly premultiplied be a correction factor based on a censum underenumeration estimate (the details of this premultiphation are given with the derivation of the preliminary correction lactors in Appendix A); (2) total registered intercensal births by sex for each of the four fiverear intercensal periods, given in Tables 5 and 6 ; (3) total registered intercensal deaths by sen and five-vear dhe groups (openended at age 70) for each of the four five-vear intercensal periods, again given in Tables 5 and $\left(0 ;(4) l_{10}, l_{1}\right.$, and,$l_{1,}$ walues by see for each of the four five-vear intercensal periods obtained from Coale-Demeny model North iife tables (levels $16.825,17.175,17.525$, and 17.875 respectively for males given in Appendix Tatile B3.3, and 17.0, 17.6, 18.2, and 19.9 respectively for females given in Appendix Table B.t) in order to yiold are estimate of the separation factor for deaths at ages ( $)-4$ by sex for cach of the four five-year intercensal periods: (3) a specified sex ratio at birth of 1.158 for each of the four fiveyear intercensal feriods; and (6) estimated preliminary correction faciors
for each of the demographic quantities listed in (1), (2), and (3). Because there are fifteen age groups (fourteen five-year age groups plus the openended one), for each sex there wil: be forty-five census correction factors (fifteen for each census) and sixty ceath correction factors: (fifteen for each five-year intercensal period). If we add to this the four correction factors for births for each sex (one for each five-year intercensal peric d), we have a total of 109 correction factors for each sex. The preliminary correction factors are put into the computer program with male factors first and female factors following. The order of preliminary correction factors for each sex is births first (in chronological order), then censuses (youngest to oldest age group in first census, in second census, etc.), then deaths (youngest to oldest age group in first interiensal period, in second intercensal period, etc.).

Because registered births and deaths (by age) for Thailand are tabulated by calendar year, it was necessary to interpolate to obtain registered births and deaths for the four five-year intercensal periods. For example, registered births for the first five-year intercensal period, 1 April 1960 to 1 April 1965, include three-fourths of the births for 1960 , all of the births for 1961-64, and one-fourth of the births fo. 1965. The same interpolation procedure was used for deaths by age group.

Coale-Demeny model North life tables have been used frequently in the analysis of Thai data, for example by Pejaranonda Arnold, and Choe (1985) and by Fulton (1979) after a careful study comparing the fit of the different Coale-Demeny (1966) moded life table families to Thai mortality experience. We opted for such model tables rather than those resulting from the Surveys of Population Change (SPC) of 1964-67 and 1974-76 (Thailand, National Statistical (Office 1969, 1978) because of the irregularity of some SPC in, values, probably due to the small sample size. At the mortality levels listed in (4) above, the Coale-Demeny model North tables are applied both for estimating separation factors for deaths at ages (0-4 and for deriving the preliminary correction factors for deaths. The one exception involves the later adult ages, where it was judged that the spe life taioles more accurately reflect Thai mortality experience during 19n()-80); therefore the SPC life tables were employed for both sexes rather than the model North life tables in order to derive the preliminary correction factors for deaths for the $70+$ age group for cach of the four five-year intercensal periods.

Moreover, the life expectancy values ( $e_{0}$ ) from the spe life tables determined the levels of the model North life tables that were used. For example, the 1964-67 and 1974-76 SPC male life tables list ', values of 56.27 and
58.00 respectively. These values were assumed for the dates 1 April 1965 and 1 April 1975, which are encompassed by the respective SPCs. Model North levole of 17.0 and 17.7 were matched to these male $t_{0}$ values at these Ahes; accordingly, life tables for males for each of the four five-year intercensal periods were obtained by linear interpolation, taking the reference dates to be the midpoints of the five-vear periods.

The same process was used to ohtain a temale life table for each fiveyear intercensal period, exceot that level 14 was used for the period I April 1975 to 1 April 198i) rather than 18.8 as would result from linear interpolation. The decision to use this higher level was based on evidence that life expectancy for females was accelerating in the .970 s, at least in relation to that for males (Hill 1979). Moreover, the results from using the lower figure of 18.8 contradicted preliminary evidence from the 198.4-86 SPC that death registration completeness leveled off in the 1970 s rather than continuing its carlier decline.

The preliminary correction factors for deaths by sex and by five-year intercensal period for the $70+$ age group were obtained from the SPC life tables as follows. The SPC life tables for 1964-67, and 1974-76 were ascribed to the dates 1 April 1965 and 1 April 1975 respectively Linear interpolation was performed, and agein the reference dates were taken to be the midpoints of the five-vear intercensal intervals.

The estimate of 1.058 as the true sex ration at birth was based on the average of the values 1.055 and 1.062 obtained by the 1964-67 and 1975-76 SPCs. The same value of 1.0 . 8 was input for each of the five-vear intercensol periods since there is no reason to expect the sex ratio at birth to vary from one five-year peried to the next. The correction procedure then requires that the corrected registered births satisfy the specified sex ratio of 1.058 for each fise-fear internensal period. This teature has been added to the correction procedure sime its original description in Luther and Retherford (1986a, l98ob).

The complete details of the derivation of the preliminary correction factors are given in Appendis A. The reader who is interested in the mechanics, as well as the results, of the consistent correction procedure can refer to that appendix. As mentioned before, the more theoretical aspects of the methodology of the procedure are given in lather and Retherford 1986a, 1986b).

Finally, it is worth noting that the cost of running the computer program for this application involving three That censuses and four five-vear intercensil periods is only about $\$ 16$.

## PRINCIPAL FINDINGS

The principal findings are given in Tables 1-11, Figures 1-5, and Appendix Tables B.1-B.5.

## Censuses

Tables 1-4 and Figures 1 and 2 present the final results for the censuses. In Figures 1 and 2 the final correction factors of Table 3 are compared with the corresponding preliminary correction factors given in Appendix Table B. 2 .

Especially striking is the high degree of underenumeration of the 0-4 age group for both sexes for all censuses, hut especially for the 1960 census (Table 4). This is consistent with the observation of Chamatrithirong, Debavalya, and Knodel (1978), Vanel on Thailand (1980), United Nations (1966), and others that the ()-4 age group was especially underenumerated in the 1960 census because of age rounding owing to the fact that the census question asked for ages of household members rather than dates of births. Thus the ages of children between exact ages $4 \frac{1}{2}$ and 5 were frequently rounded upward to age 5 . This was less likely to occur in the 1970 and 1980 censuses, in which date of birth was asked rather than age.

If there was rounding to the nearest age at all ages in the 1960 census, then the count of, soy, the $5-4$ age group (exact ages 5 to ! 0 ) induded some children between exact ages $41 / 2$ and 5 and excluded some between exact ages $91 / 2$ and 10 . Suppose the extent of rounding was substantial and of similar proportions throughout the entire population. Since the total population between exact ages $41 / 2$ and 5 exceeds that between exact ages $91 / 2$ and 10 , such rounding would tend to result in a relative overcount of the 5-9 age group and, for the same reason, those age groups above it. Thus for those age spans in which enumeration was reiatively complete, slight overenumeration would appear. The results show this to be true for ages 20-49 for both males and lemales in the 1960 consus (see Table 4). Since the 1960 census is believed to have given a relatively complete count for ages 5 and above (Linited Nations [966), such slight overemumeration for some age groups is plausible.

On the other hand, the results imply that the 1970 census contained considerable underenumeration at ages $20-49$ as well as at ages $5(0-69$. ()f course, the total population of the 1970 census was substantially more underenumerated than the $19(0)$ census (see, e.g., Fulton 1974; Preston and Hill 1980; and Luther 1983 as well as Table 4).

Table $t$ also indicates considerabie age exageperation at the older ages in both the 1970 and 1990 censuses, although the dge exaggeration for males in 1970 was not sufficient to offee the general underount. One cannot discern from the procedure, howeser, whether such dge exaggeration occurred

Table 1. Reported and final corrected census age distributions and estimates of completeness for males: Thailand, censuses of 1960,1970, and 1980

| Age group | 1460 |  | 1970 |  | 1980 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Reported | Currected | Reporied | Correited | Reported | Corrected |
| $0-4$ | $2,141.554$ | $2.512,313$ | 2,860.Fャ2 | 3,234,784 | 2,771.779 | 3,040,599 |
| 5-9 | 2.016 .266 | 2.011 .074 | 2,682, 2 601 | 2,757,912 | 2,979,485 | 3,204,070 |
| 10-14 | 1, Э¢5, isc | 1,558, 905 | 2.312 .473 | 2,353,701 | 3,006,300) | 3,071,068 |
| 15-19 | 1,265, 152 | 1.297.945 | 1.834,446 | 1.932,295 | 2,696,618 | 2,676,213 |
| 20-24 | 1.214 .350 | 1.202,446 | 1,323,314 | $1.300,287$ | 2,239,837 | 2,275,485 |
| 25-29 | 1,026,654 | 1,014,704 | 1.099,47.3 | 1,236,913 | 1,743,323 | 1,847,351 |
| 30-34 | 885,681 | 8 8.6.690 | 1,048, 6.44 | 1,141,353 | 1,333,155 | 1,423,980 |
| 35-39 | 643.448 | 690.219 | 954.165 | 960,952 | 1,161,496 | 1,170,734 |
| 40-4 | 569.744 | 563.445 | 775.308 | 818.919 | 1,064,541 | 1,075,500 |
| 45-49 | 494.691 | 487.183 | 599.876 | 034.543 | 927,227 | 891,716 |
| 50-54 | 402.425 | 403.137 | 472,783 | 507,599 | 74.4.588 | 744,492 |
| 55-59 | 322.258 | 322,151 | 388,820 | 424,953 | 543,743 | 555,018 |
| 60-64 | 229,018 | 242.226 | 301.182 | 333,743 | 411,260 | 422.898 |
| 65-69 | 149,291 | 157,901 | 213,227 | 245,685 | 296,774 | 328,157 |
| $70+$ | 177,685 | 187.933 | 250,973 | 267,276 | 408,481 | 395,836 |
| Total | 13,154,122 | 13,522.828 | 17,123,86! | 18350,923 | 22,328,607 | 23,123,208 |
| Completeness (\%) | 97.27 |  | 93.31 |  | 96.56 |  |

Sources: Reported figures are from Thailand, Central Statistical Office (1962: table 3, p. 9); Thailand, National Statistical Office (1972: table 4. p. 12; 1983: table 4, pp. 22-23). Unknown age's are allocated proportionally.

Table 2. Reported and final corrected census age distributiors and estimates of completeness for females: Thailand, censuses of 1960, 1970, and 1980

| Age group | 1960 |  | 1970 |  | 1980 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Reported | Corrected | Reported | Corrected | Reported | Correcied |
| 0-4 | 2,105,216 | 2,439,170 | 2.709,770 | 3,117,372 | 2,654,067 | 2,954,623 |
| 5-9 | 1,982,900 | 1,974,806 | 2,609,020 | 2,673,382 | 2,955,893 | 3,129,504 |
| 10-14 | 1,527,751) | 1,520,03¢ | 2,255,500 | 2,308,528 | 2,898,19; | 3,000,334 |
| 15-19 | 1,238,223 | 1,242,389 | 1,887,756 | 1,918,301 | 2,711,64, | 2,616,050 |
| 20-24 | 1.206,022 | 1.198.094 | 1,363,440 | 1,475,896 | 2,281,192 | 2,257,910 |
| 25-29 | 1,048,097 | 1,038,303 | 1,14, 824 | 1,200,251 | 1,811,104 | 1,865,258 |
| 30-34 | 871.233 | 867,134 | 1,078,451 | 1,154,181 | 1,365,493 | 1,426,153 |
| 35-39 | 681,001 | 677,001 | 958,819 | 992,093 | 1,183,324 | 1,155,458 |
| 40-44 | $56+.692$ | 561,835 | 767,302 | 819,472 | 1,102,535 | 1,107,333 |
| 45-49 | 483,720 | 480,027 | 598,210 | 633,529 | 967,187 | 940,695 |
| 50-54 | 410.99 .4 | +18,251 | 490,414 | 519,371 | 768,793 | 765,667 |
| 55-59 | 329,554 | 334,693 | 402,239 | 434,036 | 567,740 | 577,231 |
| 60-64 | 245,371 | 262,045 | 324,633 | 365.000 | +42,596 | 457,181 |
| 65-69 | 163,855 | 174,990 | 239.203 | 271,841 | 333,312 | 358,791 |
| $70+$ | 245,099 | 261,755 | 353,933 | 347.895 | 552,857 | 522,357 |
| Total | 13,103,737 | 13,448,530 | 17,273,514 | 18,231,149 | 22,495,933 | 23,134,546 |
| Completeness (\%) | 97.44 |  | 94.75 |  | 97.24 |  |

[^0]Tahle 3. Final correction factors: Thailand, censuses of 1960, 1970, and 1980

| Age group | 1960 |  | 1970 |  | 1980 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Mates | Females | Males | Femalos |
| $0-4$ | 1.173 | 1.158 | 1.129 | 1.113 | 1.097 | 1.113 |
| 5-9 | . 992 | 99\% | 1.028 | 1.025 | 1.075 | 1.096 |
| 10-14 | . 495 | . 995 | 1.018 | 1.024 | 1.021 | 1.035 |
| 15.19 | 1.026 | 1.003 | 1.053 | 1.016 | . 992 | . 965 |
| 20-2.4 | .990) | . 993 | 1.134 | 1.083 | 1.016 | . 990 |
| 25-29 | . 993 | . 491 | 1.125 | 1.049 | 1.060 | 1.030 |
| 30-34 | .989 | (9)3 | 1.088 | 1.070 | 1.068 | 1.044 |
| 35-39 | . 995 | . 994 | 1.007 | 1.034 | 1.008 | . 977 |
| 40-4 | . 990 | . 995 | 1.056 | 1.068 | 1.010 | 1.005 |
| 45-49 | . 985 | . 992 | 1.058 | 1.059 | . 962 | . 973 |
| 50-54 | 1.002 | 1.018 | 1.073 | 1.059 | 1.000 | . 996 |
| 55-59 | . 949 | 1.016 | 1.092 | 1.079 | 1.020 | 1.017 |
| 6 ()-64 | 1.058 | 1.068 | 1.109 | 1.125 | 1.029 | 1.033 |
| 65-69 | 1.058 | 1.068 | 1.152 | 1.136 | 1.105 | 1.076 |
| $70+$ | 1.058 | 1.068 | 1.065 | . 983 | . 969 | . 945 |

Nite: The entries here may be whtai ned by dividing the corrected population by the corresponding reporsed population in Tablers 1 and 2 .
in the 1960 census. With the choice of the same preliminary correction factors for the last three age groups in the fiist of the three ce nsuses and the use of the "proportional weights," mentioned earlier, to spread the correction evenly, it can be shown that the procedure will necessarily yield the same final correction factors for those last three age groups in the first census. thise the figures 1.058 for males and 1.068 for females for the 1960 census should be considered as the final correction factors (and, correspondingly, 94.5 percent for males and 93.6 percent for females as the final completeness estimates) for the combined $6(0)+$ age group rather than for the $6(0)-64,65-69$, and $70+$ age groups separately.

As; Figures 1 and 2 indicate, the preliminary and final correction factors for the censuses agree quite well except at the late ages. This shows that a considerable degree of consistemy is already present with the determination of the preliminary correction factors (see Appendix A).

## Vital registration

The vital registration results are given in Tables 5-8 and in Figures 3 and 4 , in which preliminary and final correction factors for deaths are compared.

Table 4. Completeness estimates (percentages): Thailand, censuses of 1960, 1970, and 1980

| Age group | 1960 |  | 1970) |  | 1980 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females | Males | Females |
| 0-4 | 85.3 | 86.4 | 88.6 | 89.8 | 91.2 | 89.8 |
| 5-9 | 100.8 | 100.4 | 97.3 | 97.6 | 93.0 | 91.2 |
| 10-14 | 100.5 | 100.5 | 98.2 | ¢7.7 | 97.9 | 96.6 |
| 15-19 | 97.5 | 99.7 | 95.1 | 98.4 | 100.8 | 103.6 |
| 20-24 | 101.0 | 100.7 | 88.2 | 92.3 | 98.4 | 101.0 |
| 25-29 | 100.7 | 100.9 | 88.9 | 45.3 | 94.3 | 97.1 |
| 30-34 | 101.1 | 100.7 | 91.9 | 93.5 | 93.6 | 95.8 |
| 35-39 | 100.5 | 100.6 | $99 . .7$ | 96.7 | 99.2 | 102.4 |
| 40-44 | 101.0 | 100.5 | 94.7 | 93.6 | 99.0 | (9). 5 |
| 45-49 | 101.5 | 100.8 | 94.5 | 94.4 | 104.0 | 102.8 |
| 50-54 | 99.8 | 98.2 | 93.2 | 94.4 | 100.0 | 100.4 |
| 55-59 | 100.1 | 98.4 | 91.6 | 42.i | 98.0 | 98.3 |
| 60-64 | 94.5 | 43.6 | 90.2 | 88.9 | 97.2 | 96.8 |
| 65-69 | 94.5 | 93.6 | 86.8 | 88.11 | 90.5 | 92.4 |
| $70+$ | 94.5 | 93.6 | 93.9 | 101.7 | 103.2 | 105.8 |

Note: The entries here are the reciprocals (expressed an percentages) of the entries of lable 3. Alternatively, thev may be obtained by dividing the reported population be the corresponding corrected population in lablea 1 and?

The birth registration results are in reasonable agreement with those of the United Nations (1976), U.S. Bureau of the Census (1978), Hill (1979), and others. (See table 7 of Panel on Thailand 1980 for a good summary of estimates of birth registration completeness prior to 1976.) The upward trend, as well as the level, of the estimates of the completeness of birth registration is also consistent with a preliminary birth completeness estimate of about 89 percent for the combined seves from the first vear of the 1984-86 SPC (subject to change).

The death registration results also compare quite well with other estimates. (Again, table 7 of Pand on Thailand 1980 gives a good summary of estimates of death registration completeness.) The results indicate a decrease in death registration completeness during the 1960 s and early 1970 s and a slight upturn in the late 1970 s on the whole-an upturn for males and a leveling off for females (Tables 5 and 6 ). This trend and the level are consistent with the results of the 1964-67 and 1974-76 SPCs and a preliminary finding of the first year of the $1984-86 \mathrm{SPC}$. which gives an estimate of about 76 percent completeness for death registration of both sexes combined (subject to change).

Figure 1. Preliminary and final correction factors for males: Thailand, censuses of 1960, 1970, and 1980


Note: Remult are plotled at the midponts of live-vear age groups.

It is not surprising that the results indicate greatest underregistration of deaths at the youngest and oldest ages, and especially at ages 0-4 (Table 8). It is somewhat surprising, however, that death registration completeness at ages (0-4 declined steadily and substantially during the period 196()-80. This trend is present in the preliminary correction factors as well.

Another interesting leature of the results, especially evident from inspection of Figure 3 , is the high degree of underregistration of male deaths at ages 20-29 for all five-vear intercensal periods. This trend also emerges for females (Figure 4 ), but only after 1970 ).

The preliminary and final correction factors for deaths show considerable agreement except at the late ages, as was true for the censuses (Figures 3 and 4). Thus the preliminary correction factors as a whole reflect a considerable degree of consistency, although less so at the late ages.

Figure 2. Preliminary and final correction factors for females: Thailand, censuses of 1960, 1970, and 1980


Note: Results are plotted at the midpuints of five-vear age groups.

## Life tables

The intercensal life tables for the four five-year intercensal periods, which are derived from the results of the consistent correction procedure, are given in Tables 9 and 10, and the life table death probabilities for the extreme periods 1960-65 and 1975-80 are shown in Figure 5. The life tables are computed directly from corrected deaths and population by age, as detaled in Appendix C where precise formulas are given. The calculations use the estimated midcensus age-sex distributions of Table 11 produced by the consistent correction procedure, which pertain to 1 April 1965 and 1975, and the corrected data given in Tables 1, 2, 5, and 6 .

Table 5．Registered and final corrected intercensal births and deaths and estimates of registration completeness for males：Thailand，1960－65，196ラ－70，1970－75，and 1975－80

| Births deaths and age group | 14001－65 |  | 1465－70 |  | 1970－75 |  | 1975－80 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Repurted | Cirrected | Reported | Cirrected | Reported | Corrected | Reported | Corrected |
| Births | 26.30 .151 | 3，194．78 | 2.944 .43 | 3，562，74n | 3，050，797 | 3，628．987 | 2，811，543 | 3，318，567 |
| Completeness（ ${ }^{\circ} \mathrm{O}$ ） | \＄2．33 |  | 82.65 |  | 84.17 |  | 84.72 |  |
| Deaths |  |  |  |  |  |  |  |  |
| 0－4 | 224.812 | 400， 58. | 191.302 | ＋14．457 | 161．667 | 408.795 | 114.808 | 353，137 |
| 5－9 | 33，20\％ | 53，298 | 32，834 | 56.786 | 30， 189 | 60,236 | 26.830 | 58，212 |
| 10－14 | 12.047 | 24.933 | 18.345 | 27，943 | 18，979 | 30，080 | 18，725 | 32，687 |
| 15－19 | 15，スマッ | 26．229 | 14.579 | 31，394 | 24，350 | 36，402 | 28.405 | 40,395 |
| 20－24 | 17.40 n | 3i，34， | 14.474 | 34，681 | 25.213 | ＋2， 108 | 34，247 | 49，555 |
| 25－29 | 17.153 | 30.468 | 18，667 | 30,907 | 21，26i | 33，789 | 28，100 | 41，822 |
| 30－34 | 18.542 | 27，6．38 | 20.938 | 31， 414 | 22，404 | 31，309 | 25，424 | 34，895 |
| 35－39 | 20.504 | 20， 83.5 | 23，819 | 29．438 | 20.012 | 32，171 | 28，646 | 32，365 |
| 40－4 | 21.800 | 25.428 | 25.539 | 30， 134 | 29，430 | 32，950 | 34，024 | 35，861 |
| 45－49 | 25．87\％ | 25.808 | 27．100 | 28，457 | 31，377 | 34，913 | 37，72．4 | 37，489 |
| 50－54 | 29.031 | 30，362 | 31.7211 | 32,419 | 33，649 | 37，130 | 39，701 | 43，501 |
| 55－59 | 30.562 | 31,804 | 34.254 | 35.791 | 36，937 | 39，127 | 39，933 | 43，506 |
| 60－64 | 32.972 | 35，684 | 38，825 | 40，830 | ＋2，805 | 47，027 | 47，810 | 49，641 |
| 65－69 | 29，125 | 35.565 | 36.553 | $4+6012$ | i2，954 | 51，940 | 47，423 | 57，736 |
| 71）＋ | 70.276 | 106.767 | 96,868 | 138，106 | 116，912 | 164，488 | 135，619 | 182，003 |
| Total | 610，129 | $415,10 \mathrm{n} 3$ | 635，817 | 1，011，375 | 6n＋4， 139 | 1，082，462 | 687，419 | 1，092，807 |
| Completeness（\％） | 6n．to |  | 62.87 |  | 61.35 |  | 62.90 |  |

Sources：Registered births：Thailand，National Statistical Office（n．d．：No．30，1972－73，table 29，p．84；n．d．：No．32，1976－80，table 23，p．45； 1982）．Registered deaths：Thailand，Ministry of Public Health，Division of Vital Statistics（1962，1965，1968，1972，1977，1983）．

Table 6. Registered and final corrected intercensal births and deaths and estimates of registration completeness for females: Thailand, 1960-65, 1965-70, 1970-75, and 1975-80

| Birthsideaths and age group | 1960-63 |  | 1965-70 |  | 1970-75 |  | 1975-80 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Reported | Corrected | Reported | Corrected | Keported | Corrected | Reported | Corrected |
| Births | 2,363,132 | 3,019.648 | 2,715,779 | 3,367,435 | 2,855,596 | 3,430,045 | 2,668,218 | 3,136,641 |
| Completeness (\%) | 78.2n |  | 80.65 |  | 83.25 |  | 85.07 |  |
| Deaths |  |  |  |  |  |  |  |  |
| 0-4 | 179.14.4 | 326.591 | 156.281 | 322,376 | 128,101 | 299,387 | 87,644 | 230,302 |
| 5-9 | 28.697 | +4.300 | 28.574 | +4.230 | 26,629 | 2, $+3,815$ | 22,848 | 230,302 36,716 |
| 10-14 | 13,066 | 22.003 | 14,417 | 23,249 | 15,082 | 23,511 | 14,225 | 22,567 |
| 15-19 | 12.699 | 20,210 | 14,951 | 23,460 | 17,026 | 25,13; | 17,742 | 24,770 |
| 20-24 2-29 | 17,128 | 20,309 | 16,315 | 22,608 | 17,126 | 26,552 | 17,607 | 27,821 |
| 25-29 $30-34$ | 17.818 18.996 | 22,999 20.721 | 16,064 18,593 | 21.149 | 14,373 | 23,553 | 14,539 | 26,578 |
| $30-34$ $35-39$ | 18,996 19.349 | 22.721 21.737 | 18,593 20,059 | 23,368 23,332 | 16,417 19,252 | 21,450 | 14,281 | 22,803 |
| 40-4 | 18,638 | 21.737 20.457 | 20,059 <br> 19886 <br> 18.83 | 23,332 23,534 | 19,252 21,014 | 24,170 25,873 | 16,822 21,218 | 21,778 26,299 |
| +5-49 | 12.435 | 19,517 | 19.035 | 21,216 | 20,562 | 24,908 | 21,218 | 26,299 |
| 50-54 | 15.896 | 22,208 | 21,055 | 23,738 | 22,454 | 26,417 | 24,628 | 30,375 |
| 55-59 | 20,050 | 24,879 | 22,003 | 26.519 | 24,353 | 28,759 | 25,672 | 30,896 |
| 60-64 | 22, $1+4$ | 29.241 | 27.055 | 32,896 | 28,6\%3 | 35,524 | 31,768 | 38,309 |
| 65-69 | 21,484 | 32,793 | 26,940 | 38,687 | 31,284 | 43,940 | 33,738 | 47,897 |
| $70+$ | 80,530 | 128,931 | 103,983 | 155,187 | 122,355 | 171,486 | 140,629 | 205,244 |
| Total | 506,628 | 778,916 | 525,211 | 825.548 | 524,701 | 844,480 | 505,662 | 818,809 |
| Completeness (\%) | 05.04 |  | 63.62 |  | 62.13 |  | 61.88 |  |

Sources: Same as Table 5.

Table 7. Final correction factors for registered births and deaths: Thailand, 1960-65, 1965-70, 1970-75, and 1975-80

| Births deaths and age group | 19n(1)-65 |  | 1965-70 |  | 1970-75 |  | 1975-80 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Male | Femaler | Aates | Females | Males | Fencaies | Males | Females |
| Births | 1.215 | 1.274 | 1.210 | 1.240 | 1.190 | 1.201 | 1.180 | 1.176 |
| Deaths |  |  |  |  |  |  |  |  |
| ()-4 | 1.809 | 1.824 | 2193 | 2.063 | 2.529 | 2.337 | 3.076 | 2.628 |
| 5-9 | 1.605 | $1.5+4$ | 1.724 | 1.548 | 1.995 | 1.645 | 2.170 | 1.607 |
| 10-14 | i.th3 | 1.610 | 1.523 | 1.613 | 1.585 | 1.559 | 1.746 | 1.586 |
| 15-19 | 1.068 | 1.591 | 1.603 | 1.569 | 1.495 | 1.476 | 1.422 | 1.396 |
| 20-24 | 1.735 | 1.186 | 1.781 | 1.386 | 1.670 | 1.550 | 1.447 | 1.580 |
| 25-29 | 1.77 | 1.291 | 1.656 | 1.317 | 1.589 | 1.639 | 1.488 | 1.828 |
| 30-34 | 1.491 | 1.196 | 1.454 | 1.257 | 1.397 | 1.307 | 1.373 | 1.597 |
| 35-39 | 1.309 | 1.121 | 1.215 | 1.163 | 1.237 | 1.253 | 1.130 | 1.295 |
| 40-44 | 1.166 | 1.098 | 1.180 | 1.183 | 1.120 | 1.231 | 1.054 | 1.239 |
| 45-49 | 1.001 | 1.119 | 1.069 | 1.115 | 1.113 | 1.211 | . 994 | 1.135 |
| 50-54 | 1.046 | 1.175 | 1.022 | 1.127 | 1.103 | 1.177 | 1.096 | 1.233 |
| 55-54 | 194 | 124 ! | 1.045 | 1.205 | 1.059 | 1.181 | 1.089 | 1.203 |
| O0-64 | 1.082 | 1.320 | 1.052 | 1.216 | 1.099 | 1.239 | 1.038 | 1.206 |
| 65-é | 1.228 | 1.526 | 1.220 | 1.436 | 1.209 | 1.405 | 1.217 | 1.420 |
| $70+$ | 1.4 $\%$ | 1.601 | 1.426 | 1.492 | 1.407 | 1.402 | 1.342 | 1.459 |

[^1]Table 8. Completeness estimates (percentages) for death registration: Thailand, 1960-65, 1965-70, 1970-75, and

| Age group | 1960-65 |  | 1965-70 |  | 1970-75 |  | 1975-80 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females | Males | Females | Males | Females |
| 0-4 | 55.3 | 54.8 | 45.6 | 48.5 | 39.5 | 42.8 | 32.5 |  |
| 5-9 | 62.3 | \%. 8 | 57.8 | 64.6 | 50.1 | 60.8 | 32.5 46.1 | 38.1 62.2 |
| 10-14 | 68.4 | 62.1 | 65.7 | 62.0 | 63.1 | 64.1 | 57.3 | 62.2 63.1 |
| 15-19 | 60.0 | 62.9 | 62.4 | 63.7 | 66.9 | 67.8 | 70.3 | 71.6 |
| 20-24 | 57.6 | 84.3 | 56.1 | 72.2 | 59.9 | 64.5 | 69.1 | 63.3 |
| 25-29 | 56.3 | 72.5 | 60.4 | 75.9 | 62.9 | 51.0 | 67.2 | 54.7 |
| 30-34 | 67.1 | 83.6 | 68.8 | 79.6 | 71.6 | 76.5 | 72.8 | 62.6 |
| 35-39 | 76.4 | 89.2 | 82.3 | 86.0 | 80.8 | 79.7 | 88.5 | 77.2 |
| $40-44$ $+5-49$ | 85.8 | 91.1 | 84.7 | 84.5 | 89.3 | 81.2 | 94.9 | 80.7 |
| $45-49$ $50-54$ | 99.9 95.6 | 89.4 85.1 | 93.5 | 89.7 | 89.8 | 82.6 | 10.96 | 88.1 |
| 50-54 $55-59$ | 95.6 96.1 | 85.1 80.6 | 97.8 95.7 | 88.7 | 90.7 | 85.0 | 91.2 | 81.1 |
| 60-64 | 92.4 | 80.6 75.8 | 95.7 95.1 | 83.0 82.2 | 94.4 91.0 | 84.7 80.7 | 91.8 | 83.1 |
| 65-69 | 81.4 | 65.5 | 82.0 | 69.6 | 91.0 82.7 | 80.7 71.2 | 96.3 82.2 | 82.9 70.4 |
| $70+$ | 71.4 | 62.5 | 70.1 | 67.0 | 71.1 | 71.3 | 74.5 | 70.4 68.5 |

Note: The entries here are the reciprocals (expressed as percentages) of the entries of Table 7 . Alternativelv, they may be obtained by dividing
the reported deaths by the correspondiny corected deaths in the reported deaths by the corresponding corrected deaths in Tables 5 and 6.

Figure 3. Preliminary and final correction factors for intercensal registered deaths for males: Thailand, 1960-65, 1965-70, 1970-75, and 1975-80


Notes: Results are ploted at the midpoints at five-year age proups For births, the preliminary and final correction factors are: 1960-65, 1.224 and $1.215 ; 1965-70,1.220$ and 1.210 ; $1970-75,1.199$ and 1.190 ; 1975.80, 1.179 and 1.180).

The life tables generally show morality improvements with time, as one would expect. The improvements are more substantial for females than males (Figure 5). The improvement in life expectancy is greatest for the youngest age group, in accordance with the pattern in most countries. Both the $5 \%_{0}$ estimates, based on marked underregistration of deaths at ages (0-4, and the life expectancy estimates agree quite well with those given by the SPCs for 1964-67 and 1974-76, and by Rungpitarangsi (1974) for 1960 and 1970 ). Moreover, the widening gap between male and female life expectancies during the period $196(0)-80$ is consistent with the conclusions of Hill (1979:34-35) and the results of the SPCs.

Figure 4. Preliminary and final correction factors for intercensal registered deaths for females: Thailand, 1960-65, 1965-70, 1970-75, and 1975-80


Notes: Results are plotted at the midpoints of five vear dge groups. Fire birthes, the preliminary and final correction faturs are: Wat 65, 1.300 and 1278 ; 1065 -70. 1.250 and 1.240 ; $1970-75,1.211$ and 1.201 : 1475 80. 1.174 and 1.176 .

Finally, comparison of the $196(1-65$ and $1965-70$ male life tables reveals a possible slight anomaly in the results. The $T_{01}$ value decreases a bit whereas the $l_{70}$ volue increases. This medns that proportionately more people survived t: age 70 according to the $1965-70$ life table than according to the 1990-65 one, but that the greater number lived fewer person-vears thereafter. This reversal may be real. Another possibility is that the SPC life tables do not reflect the mortality experience of males accurately at the oldest ages. This reversal is slight, however, compared with a more pronounced reversal of this same type that results for females when one uses model North life tables rather than the SPC life lables to determine the prelimi-

Figure 5. Corrected life table probabilities of dying, $5 q_{x}$ : Thailand, 1960-65 and 1975-80


nary correction factors for the $70+$ age group. It was primarily on this basis that the SPC life tables were chosen to determine the preliminary correction factors for the $70+$ age group.

## SUMMARY AND (ONCLUSION:

This paper has appled a new procedure tor the simulaneous and consistent correction of two or more censuses and intercensal registered births and deaths to the three That emsuses of 1960, 1970, and 1980, and to the registered birtbs and deaths for the intervening period 1960-80. The results reveal marked underenumeration of the 0 tage group in each census and especially in the logocensus, in which age monding was apparently prevalent. There is also an indialion of age exagieration at the aldest ages at least in the 1970 and 1980 censuses.

The results also indiate that birth registration mompleteness rose during 196()-80, whereds, death registration completeness dedined slightly until the early 1970 sand then lesoled off or began to rise. The greatest underregictation of deathe was at ages ( 0 - and the registration completeness for deathes at this age appears to have decreased throughout the period.

Life expectance rose for both males and females, at an average of about one-fith vear per yoar for males and two-fifths year per vear for females between 1960 and jus0) Thus bemale superiority in life expectancy inereased during the period, hom about 4.5 gears in $1960-65$ to dhout seven veats in 1975-80.

Finally, it should be noted that these results are reasonably consistent with previous estimates. The results of this paper, however, have the advantage of providing consistent corrections for multiple sets of data involving more censuses and intercensal periods than in previous studies.

Table 9. Derived intercensal life tables for males: Thailand, 1960-65,

| Age ${ }_{\text {d }}$ | 1960-65 |  |  | 1965-70 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 54. | 18 | $I_{1}$ |  | $1{ }_{s}$ | $L_{1}$ |
| 0 | . 13134 | 100, 1000 | 450,391 | 120以上 | 100,000 | 453,974 |
| 5 | . 02417 | 80,800 | +29.551 | . 02192 | 87.908 | + $3.35,061$ |
| 10 | . 01426 | 84,766 | 421.15.4 | . 01298 | 85.981 | 427,369 |
| 15 | .01850) | 83,557 | +14,24.3 | . 01821 | 84,865 | 420,906 |
| 20 | . 02428 | 82,01? | 1015,215 | . 02493 | 83,319 | 411,837 |
| 25 | . 127 (6) | 80,020 | 34.460 .4 | . 02536 | 81,243 | 401,200 |
| 30 | .122932 | 77.812 | . 383.707 | . 02832 | 79,182 | 390,700 |
| 35 | . 134770 | 75,530 | 371.773 | . 03163 | 76,434 | 378,991 |
| 40 | .) 4091 | 72.9019 | 357,648 | . 0403.4 | 74,506 | 365,798 |
| 45 | . 04945 | (6) 927 | 341,421 | . 04861 | 71,500 | 349,523 |
| 50 | .06846 | (6). $46{ }^{(1)}$ | 321,707 | .106514 | 618.024 | 329,600 |
| 55 | .14883, | 61,918 | 296,\%)2 | . $0 \times 6$ 6, $\mathrm{T}_{7}$ | 63,593 | 305,116 |
| 60 | 12764 | $56,44^{4}$ | 265,801 | . 12435 | 58,088 | 273,699 |
| 65 | 18392 | 4) 2.24 | 226,597 | . 18223 | 50,865 | 233,161 |
| 70 | 1.000\% | 41, 187 | 402,04, 3 | 1.100000 | 41,59\% | 388,560 |
| $t_{0}$ |  |  | $5+4.83$ |  |  | 55.65 |

Table 10. Derived intereensal lite tablen for females: Thailand, 1960-65,

| $A \mathrm{Se}_{2}$ | 1960 |  |  | 1965-70 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 4 | 1 | 1. | $\cdots$ | 1. | , 1 |
| 0 | . 11184 | 1000000 | 458,300 | (1)9845 | 100.000 | 462,870 |
| 5 | . 1221.48 | 88.851 | $4.40,093$ | 01757 | 90,155 | 447.179 |
| 10 | . 01283 | 87.031 | 4.32,705 | 01097 | 88.572 | $440,6.39$ |
| 15 | . 01478 | 85.915 | 426.6.697 | 01.380 | 87.6101 | $4.35,339$ |
| 20 | . 01665 | 84.645 | +19,8,39 | 01678 | 86,386 | $428,6+4$ |
| 25 | . 020165 | 83.236 | +12,152 | 0101764 | 84, 4937 | 420, 975 |
| 30 | . 122405 | 81,518 | 4113,078 | (12141 | 83.438 | +13,010 |
| 35 | 02858 | 79.557 | 342,721 | 02530 | 81,652 | +103,513 |
| 40 | . 03328 | 77,283 | 380, 48.3 | 0318.4 | 79,586 | 392,296 |
| 45 | . 03766 | 74.711 | 366, 9.47 | . 035572 | 77,05? | 378,913 |
| 50 | . 04959 | 71,898 | 3511,949 | . 047761 | 74,299 | 303,195 |
| 55 | . 066658 | 68.336 | 331.2.41 | . 10.22 .4 | 70,761 | 34,3,319 |
| 60 | . 109863 | 6.3.786 | 304,4(6) | (09428 | 66, 357 | 317,479 |
| 65 | . 15376 | 57.495 | 268,423 | . 1.4460 | 60, 101 | 280, 3.42 |
| 70 | 1.100000 | 48.654 | 5388.142 | 1.000000 | 51,411 | 559,210 |
| $\mathrm{c}_{0}$ |  |  | 59.26 |  |  | 60.87 |
|  |  |  | < |  |  |  |

1965-70, 1970-75, and 1975-80

| $\mathrm{Agex}^{\prime}$ | 1970-75 |  |  | 1975-80 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $57 . x$ | $I_{x}$ | $L_{*}$ | $5 \%$ | $I_{x}$ | ${ }_{5} L_{x}$ |
| 0 | . 11320 | 100,000 | 455,822 | . 10458 | 100,000 | 453,119 |
| 5 | .02039 | 88,68.1 | 439,155 | . 01826 | 89,542 | 443,680 |
| 10 | . 01187 | 86,872 | 431.960 | . 01129 | 87.907 | 437,20) |
| 15 | . 01713 | 85,841 | 425,865 | . 01613 | 86,915 | 431,32.0 |
| 20 | . 02488 | 84.371 | +17.213 | . 02372 | 85.513 | 422,960 |
| 25 | . 02492 | 82,272 | 406,660) | . 02533 | 83,485 | 412,752 |
| 30 | . 02637 | 30,221 | 395,959 | . 122642 | 81,370 | 401,925 |
| 35 | .03079 | 78.106 | 38.4,948 | .02802 | 79.221 | 390,702 |
| 40 | . 03720 | 75,701 | 371,905 | . 03540 | 77,001 | 378,691 |
| 45 | . 04867 | 72,885 | 356.495 | . 04397 | 74,275 | 363,731 |
| 50 | . 06556 | 69,338 | 336,258 | . 06369 | 71,009 | 344.967 |
| 55 | .08426 | 64,792 | 311,000 | . 08230 | 66,486 | 319,893 |
| 60 | . 12437 | 59.337 | 279,474 | . 11683 | 61,014 | 288,159 |
| 6.5 | . 18057 | 51,957 | 238,038 | .17419 | 53,886 | 247,646 |
| 70 | 1.00000 | 12,375 | 403,076 | 1.00000 | 44,499 | 464,938 |
| $\iota^{\prime}$ |  |  | 56.54 |  |  | 58.07 |

1965-70, 1970-75, 1975-80

| Age, | 1970-75 |  |  | 1975-80 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $5 \%$ | 1 | St | 51.7 | 13 | ${ }_{5} L_{x}$ |
| 0 | . 08777 | 100,000 | 465,90.4 | . 07226 | 100.000 | 470,985 |
| 5 | . 01531 | 91.223 | 452,843 | . 01186 | 92.774 | 461,165 |
| 10 | . 00950 | 89,826 | 447,141 | .00801 | 91,674 | 456,653 |
| 15 | . 01199 | 88.973 | 442,430 | 01011 | 90,940 | 4.52,559 |
| 20 | . 01590 | 87.9016 | 436,469 | . 01343 | 90,021 | 447349 |
| 25 | . 01778 | 86,514 | 429,060 | . 01617 | 88,812 | 440,920 |
| 30 | . 01823 | 84,970 | 421,016 | . 01754 | 87,376 | 433,416 |
| 35 | . 02262 | 83.421 | 412,696 | . 01887 | 85,844 | 425,213 |
| i) | . 02878 | 81,534 | 402,287 | .1025 7 | 84,224 | 416,178 |
| 45 | . 03480 | 79, 187 | 389, 823 | . 03029 | 82, 104 | 404,827 |
| 50 | . 144616 | 76,432 | .374,0.32 | . 04345 | 79,617 | 390,341 |
| 55 | . 060160 | 72,90.4 | 354, 164 | . 05666 | 76,118 | 370,668 |
| 60 | . 08893 | 68,486 | 327.937 | . 08.596 | 71,805 | 344,588 |
| 65 | . 13898 | 62,396 | 292,240 | . 13159 | 65,633 | 307,631 |
| 70 | 1.00000 | 53,724 | 643,952 | 1.00000 | 56.996 | 701,129 |
| i] |  |  | 62.92 |  |  | 65.24 |

Table 11. Estimated midcensus age distributions for males and females: Thailand, 1965 and 1975 (1 April)

|  | 1965 |  |  | 1975 |  |
| :---: | ---: | ---: | ---: | ---: | :---: |
| Age | Males | Females |  | Vates |  |
| $0-4$ | $2,877,805$ | $2,767,810$ | $3,31,6,346$ | $3,196,146$ |  |
| $5-9$ | $2,396,066$ | $2,342,268$ | $3,116,518$ | $3,029,976$ |  |
| $10-14$ | $1,961,963$ | $1,941,655$ | $2,712,754$ | $2,639,719$ |  |
| $15-19$ | $1,533,324$ | $1,498,930$ | $2,320,461$ | $2,284,206$ |  |
| $20-24$ | $1,269,707$ | $1,222,129$ | $1,893,040$ | $1,892,458$ |  |
| $25-29$ | $1,172,027$ | $1,176,440$ | $1,462,338$ | $1,450,844$ |  |
| $30-34$ | 990,641 | $1,015,443$ | $1,204,364$ | $1,177,749$ |  |
| $35-39$ | 848,454 | 842,905 | $1,109,613$ | $1,131,371$ |  |
| $40-44$ | 664,088 | 655,904 | 928,391 | 967,072 |  |
| $45-49$ | 538,282 | 541,848 | 784,987 | 794,082 |  |
| $50-54$ | 459,053 | 459,164 | 598,521 | 607,866 |  |
| $55-59$ | 372,054 | 394,707 | 469,471 | 491,783 |  |
| $60-64$ | 288,405 | 307,633 | 381,876 | 401,894 |  |
| $65-69$ | 206,500 | 231,028 | 284,260 | 325,268 |  |
| $70+$ | 221,184 | 291,397 | 322,508 | 426,281 |  |
| Total | $15,799,553$ | $15,689,261$ | $20,897,448$ | $20,816,715$ |  |

Note: These re the estimated midcensus age-sex distributions given by the consistent correction procedure.

## APPENDIX A. DERIVATION OF THE PRELIMINARY CORRECTION FACTORS

## Births

A linear increase was assumed in the birth registration completeness of both sexes combined, from 78 percent on 1 April 1960 to 86 percent on 1 April 1980. This agrees quite well with the estimates of the U.S. Bureau of the Census (1978) and Hill (1979). We should note that the latter estimates are based on the Ministry of Public I lealth (MOPII) registration figures, whereas ours are based on the Ministry of Interior (MOI) figures, which are larger in most vears; see the corrigenda sheet of Hill (1979) for a comparison of these figures through 1974. A se ratio at birth of 1.058 for each five-year intercensal period was also used to help calculate the preliminary correction factors for births by sex.

From this assumption regarding birth registration completeness, it follows that the average birth registration completeness for the combined sexes was 79, 81, 83, and 55 percent for the successive five-year intercensal periods. The calculation of the preliminary correction factors for male and female births for the first five-year intercensal period will be shown in detail for illustrative purposes.

Let $k_{n \text { : }}$ and $k_{\text {f }}$ be the preliminary correction factors for numbers of male and female registered births $B_{m}$ and $\beta_{1}$ respectively for the five-year period. Then $k_{m} B_{m}$ and $k_{t} B_{y}$ denote the estimated true numbers of male and female births. Assuming a 1.058 sex ratio at birth, one has

$$
\begin{equation*}
k_{m} k_{m}=1.058 k_{q} B_{t} \tag{1}
\end{equation*}
$$

Assuming, moreover, 79 percent birth registration completeness for 1960-65, one may equate two different expressions of the true number of births for the combined sexes as follow's:

$$
\begin{equation*}
k_{m} R_{m}+k_{1} B_{1}=\left(R_{m}+B_{f}\right) /(.79) \tag{2}
\end{equation*}
$$

Since $B_{m}=2,630,151$ and $B_{1}=2,363,132$ are known, they may be substituted into cquations (1) and (2). These equations then represent a system of two linear equations in the two unknowns, $k_{m}$ and $k_{p}$, which may be solved simultaneously to yield the values of $k_{m}$ and $k_{t}$.

The same process may be used to find the preliminary correction factors for registered birthis for the other five-year intercensal periods. The values whained for the suceessive periods are 1.234, 1.220, 1.149, and 1.179 for males and 1.300, 1.250, 1.211, and 1.174 for females. They are recorded in Appendia Taole B. $\overline{\text { a }}$.

## Censuses

The preliminary correction factors for censuses are derived directly from the following corrected census age-sex distributions obtained from outside sources: (1) For the 1960 census, the correction given in United Nations (1966); (2) for the 1970 census, the correction of the 0-4 and 5-9 age groups in Arnold and Phananiramai ( ${ }^{9775}$ ), and of the other age groups in Fulton (1979); and (3) for the 1980 census, the correction given by Pejaranonda, Arnold, and Hauser (1983). We use the final census report (Thailand, National Statistical Office 1983) rather than the Advance Report (Thailand, National Statistical Office 1981) for the corrected female 1980 census population age $70+: 552,900$ rather than 562,900 . These "preliminarily corrected" census age-sex distributions are recorded in Appendix Table B. 1; and the resulting preliminary correction factors, obtained by dividing these preliminarily corrected populations by the corresponding reported populations from Tables 1 and 2, are given in Appendis Table B.2.

A certain amount of consistency is built into the preliminary correction factors for censuses, by the use of the census corrections above, each of which is derived in part from the preceding census correction. However, the preliminary correction factors for births are derived independently; and to an extent, the same may be said for the preliminary worrection factors for deaths (see below).

As with Pejaranonda, Aroold, and Hauser (1983), we have opted for the Fulton (1979) corrections of the 1970 census age-sex distributione rather than those of Arnold and Phananiramai (1975). The preliminarily corrected census age--sex distributions of Appendix Table B.1, which have been used for deriving the preliminary correction factors for censuses, are used to calculate total population preliminary correction factors for each sex at each census. The reciprocals of these factors in turn yield census underenumeration estimates bey sex for each census. The total population preliminary correction factors are given in the last row of Appendix Table B.2.

The reported census age-sex distributions (given in Tables 1 and 2 ) are uniformly premultiplied by these total population preliminary correction factors of Appendix Table 3.2 before being put into the computer. For example, each female five-year age group is premultiplied by 1.051 for the 1970 census and by 1.025 for the 1980 census. To compensate, cach preliminary correction factor given in Appendi, Table B. 2 must be divided accordingly before it is put into the computer. This will insure that each input census age-sex group population multiplied by the corresponding input preliminary correction factor still yields the preliminarily corrected census age-sex group population of Appendi Table B.1. Thus, for example, each preliminary correction factor for females must be divided by 1.051 for the 1970 census, and by 1.025 for the 1980 census, before being put into the
computer. In particular, the input preliminary correction factors for females of ages $0-4$ for the 1970 and 1980 censuses are 1.104/1.051 $=1.050$ and $1.110 / 1.025=1.083$ respectively. These adjustments of the reported census age-sex distributions, as well as of the preliminary correction factors for the censuses, must be put into the computer program to get appropriate census underenumeration levels in the final results (output). This is true because the compater program is written so that for the last census (the 1980 one in our case), there is equality between the total input population for bootl: sexes combined and the total output (corrected) population for both sexes combined. This equality was incorporated into the computer program to obtain more control over the census underenumeration levels in the final results.

Finally, except for the (1) 4 age group, in which underenumeration is very evident because of age rounding (see section on principal findings), all preliminary correction factors for both seves for the 1960 census were chosen to be i , on the basis of United Nations (1966). First of all, the opposing effects of age rounding partially compensate in all age groups except $0-4$ and $70+$, and the effect is relatively small on the $70+$ age group. Second, wher than for the (0-4 age group, the 1960 census count appears to be fairly complete and accurate by five-year age groups (United Nations 196(9). Moreover, this choice was made simply for lack of better evidence.

## Deaths

Coale-Demeny model North life tables were used to determine the preliminary correction factors for deaths for all age groups below 70 . The levels used were based on the $e_{0}$ values of the 1964-67 and 1974-76 S1 C life tables, as described in the section on application of our method to, Thailand, 196()-80. The life tables corresponding to these levels for males and females are given in Appendix Tables B. 3 and B. 3.4 respectively:

For the $70+$ age group the SPC life ables were used rather than model tables. That is, the preliminary correction factors for deaths for the $70+$ age group were derived from life tables obtained from the SPC life tables by an interpolation process described in the section on application of the method.

The 196()-65 and 1970)-75 preliminary correction factors for deaths were derived from the life tables be a forward survival method involving the 1960 and 1970 prciiminarily corrected census age-sex distributions. Similarly, the 1965-70 and 1975-80 factors were derived by a reverse survival method involving the 1970 and 1980 preliminarily corrected census age--sex distributions. The preliminarily corrected census age-sex distributions are given in Appendix Table B.1, and the preliminary correction factors for deaths are given in Appendix Table B.S.

More precisely, the forward survival formula used for the periods 1960-65 and 1970-75 to calculate the preliminary correction factor for deaths $\left(h_{a}\right)$ at age $a$ to $a+5$ is

$$
\begin{align*}
h_{a} & =\left(\begin{array}{c}
P_{a-5} \cdot \\
5 l_{a} \\
L_{a-5}
\end{array}\right)\binom{l_{a}-l_{a+5}}{l_{a}} / D_{a} \\
& =\begin{array}{c}
5\left(l_{a}-l_{a+5}\right) P_{a-5} \\
{ }_{5} L_{a-5} D_{a}
\end{array} \tag{3}
\end{align*}
$$

where $I_{a}$ denotes life table survivors at exact age $a ; L_{a}$ denotes life table person-years lived between exact ages $a$ and $a+5 ; D_{a}$ denotes the number of registered deaths of age $a$ to $a+5$ for the period; and $P_{a}$ denotes the preliminarily corrected population of the age group a to a +5 for the census immediately preceding the period in question (the 1960 census for the period 1960-65 and the 1970 census for the period 1970-75).

The rationale behind formula (3) is as follow's. In the first line of formula (3), the first factor in the numerator gives the number of those age $a-5$ to $a$ in the prelimimarily corrected census population at the beginning of the period who have survived to attain age a during the period, based on the life table used for the period. The second factor in the numerator is the life table probability of dying between age $a$ and $a+5$ for the period. Thus the numerator, which is the product of these tion factors, gives an estimate of the true number of deaths between ages $a$ and $a+5$ for the period. When divided by $D_{a}$, the registered number of deaths between ages $a$ and $a+5$ for the period, this yields the estimated (preliminary) correction factor for deaths for this age group.

Similarly, the reverse survival formula used for the periods 19t. 70 and 1975-80) is

$$
\begin{align*}
h_{a} & =\left(\begin{array}{cc}
l_{a} \cdot & 5 l_{a} \\
5 L_{a}
\end{array}\right)\binom{l_{a}-l_{a+5}}{l_{a}} / D_{a} \\
& =5\left(l_{a}-l_{a+5}\right) P_{a} \tag{4}
\end{align*}
$$

where the notation is as above except that $P_{10}$ is the preliminarily corrected population of the age group $a$ to $a+5$ for the census immediately following the period in question.

Formula (4) is the same as formula (3) except for the first factor in the numerator. In both formulas, this first factor estimates the true number of people who have attained age a during the period. In formula (4), the number who have attained age $a$ during the period is obtained by reverse survival of those aged a to $a+5$ in the preliminarily corrected census
population at the end of the period rather than by forward survival of those aged $a-5$ to $a$ in the preliminarily corrected census population at the beginning of the period, as in formula (3).

Clearly the forward survival formula for 1960-65 and 1970-75 is not applicable for the age group 0-4. However, its natural analogue is

$$
\begin{equation*}
h_{0}=\frac{\left(l_{0}, l_{5}\right) B}{l_{10} D_{0}} \tag{5}
\end{equation*}
$$

where $\beta$ denotes the estimated true number of births for the period and the other notation is as before. Equation (5) is used to determine the preliminary correction factor for deaths at ages $0-4 . B$ is obtained by using the preliminary correction factors for births determined above to estimate the true number of births. For example, the preliminary correction factor for deaths for females of ages $0-4$ for the period $1960-65$, given in Appendix Table B.5, is

$$
h_{11}=\begin{gathered}
(100,000-89,462)(1.300)(2,363,132) \\
(100,000)(179,043)
\end{gathered}=1.808
$$

Here, 89,462 is the $l$, value from the Coalo-Demeny model North level 17.0 female life table.

Special formulas are needed also for the preliminary correction factor for deaths for the $70+$ age group. The approximate reverse survival formula that is used may be obtained from equation (4) by letting $a=70$ and $a+$ 5 be an age beyond which anyone lives. Hence it is the same as equation (4) with $a=70$; $l_{a}+5$ set equal to 0 ; $P_{a}$ and $D_{a}$ replaced by the preliminarily corrected census population $P_{70}$, and registered deaths $D_{70}+$ respeclively; and $L_{a i}$ replaced by $T_{70}$, the life table person-years lived above age 70. Thus

$$
h_{70}=\begin{align*}
& 5 l_{70} P_{71},  \tag{6}\\
& \Gamma_{71} D_{71},
\end{align*}
$$

As before, $P_{70}$, refers to the census immediately following the period in question.

For example, the volue of $h_{70}$ for males for the period 1965-70 is

$$
h_{50}=\begin{gathered}
5(43,881)(280,900) \\
(428,383)(96,868)
\end{gathered}=1.485
$$

(Recall that the interpolated SPC life tables are used for the $70+$ age group.)
If census and death registration data and SI'C life tables for five-year age groups bevond age 70 are wailable, as is true for some periods, one may use a more precise lormula than equation (6) for $h_{70}$. This formula would consist of a sum with terms like equation (4) for the five-year age
groups and equation (6) for the open-ended age group. However, it was found that this more cumbersome formula yielded correction factors that differed little from those given by formula (6).

The reverse survival equation (6) was used to determine $h_{70}$ not only for the periods 1965-70 and 1975-80, but also for the periods 1960-65 and 1970-75 after forward surviving the age group $65+$ preliminarily corrected population of the 1960 and 1970 censuses to 1965 and 1975 respectively. It can be shown that the same result is obtained by using the forward survival formula

$$
h_{70}=\frac{5 l_{70} P_{65+}}{T_{65} D_{70}+}
$$

where $P_{65+}$ is the preliminarily corrected population of the age group $65+$ for the census immediately preceding the period in question.

## APPENDIX B. TABLES

Table B.1. Preliminarily corrected census age-sex distributions: Thailand, 1960, 1970, and 1980

| Age group | 1960 |  | 1970 |  | 1980 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females | Males | Females |
| 0-4 | 2,517,000 | 2,448,000 | 3,219,100 | 3,090,100 | 3,074,700 | 2,945,500 |
| 5-9 | 2,016,300) | 1,982,900 | 2,725,400 | 2,64,500 | 3,208,900 | 3,083,900 |
| 10-14 | 1,565,400 | 1,527,800 | 2,350,100 | 2,300,700 | 3,103,700 | 2,975,700 |
| 15-19 | 1,265,20\% | 1,238,200 | 1,904,900 | 1,919.400 | 2,672,900 | 2,587,500 |
| 20-24 | 1,214,400 | 1,206,000 | 1.475,000 | 1,480,000 | 2,283,800 | 2,265,600 |
| 25-29 | 1,026,700 | 1.048.100 | 1,263,200 | 1,192,300 | 1,866,600 | 1,862,800 |
| 30-34 | 885,700 | 871.200 | 1,147,400 | 1.153,900 | 1,447,200 | 1,421,600 |
| 35-39 | 693,400 | 681,000 | 965,700 | 990,200 | 1,178,700 | 1,172,900 |
| 40-44 | 569,700 | 564,700 | 826,100 | 821,400 | 1,063,100 | 1,105,100 |
| 45-49 | 494,700 | 483,700 | 6.37 .800 | 635,800 | 884,600 | 931,100 |
| 50-54 | 402,400 | 411,000 | 512,500 | 519,800 | 732,200 | 761,100 |
| 55-59 | 322,300 | 329,600 | 430,200 | 434,600 | 552,100 | 573,900 |
| 60-64 | 229,000 | 245,400 | 332,400 | 353,900 | 416,400 | 456,400 |
| 65-69 | 149,300 | 163,900 | 244,400 | 263,200 | 321,100 | 357,400 |
| $70+$ | 177,700 | 245,100 | 280,90 ${ }^{\prime}$ | 354,000 | 400,100 | 552.900 |
| Total | 13,529,700 | 13,446,000 | 18,315,100 | 18,159, 500 | 23,206,100 | 23,053,400 |

Sources: For 1960: United Nations (1966). For 1970: Arnold and Phananiramai (1975: table 11), for ages 0-4 and 5-9; Fulton (1979: tables 17, 18), for ages $10+$. For 1980. 「ejaranonda, Arnold, and i iauser (1983: table 11).

Table B. 2 Preliminary correction factors: Thailand, censuses of 1960, 1970, and 1980

| Age group | 1960) |  | 1970 |  | 1980 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females | Males | Females |
| 0-4 | 1.175 | 1.163 | 1.123 | 1.104 | 1.109 | 1.110 |
| 5-9 | 1.000 | 1.000 | 1.016 | 1.014 | 1.077 | 1.080 |
| 10-14 | 1.000 | 1.000 | 1.016 | 1.020 | 1.032 | 1.027 |
| 15-19 | 1.000 | 1.000 | 1.0138 | 1.017 | . 991 | . 954 |
| 20-24 | 1.000 | 1.000 | 1.115 | 1.085 | 1.020 | . 993 |
| 25-29 | 1.000 | 1.000 | 1.149 | 1.041 | 1.071 | 1.029 |
| 30-34 | 1.000 | 1.000 | 1.09 .4 | 1.070 | 1.086 | 1.041 |
| 35-39 | 1.000 | 1.000 | 1.012 | i.039 | 1.015 | . 991 |
| 40-44 | 1.000 | 1.000 | 1.066 | 1.071 | . 999 | 1.002 |
| 45-49 | 1.000 | 1.000 | 1.06 .3 | 1.063 | . 954 | . 963 |
| 50-54 | 1.000 | 1.000 | 1.06. 4 | 1.060 | . 983 | . 990 |
| 55-59 | 1.000 | 1.000 | 1.1106 | 1.080 | 1.015 | 1.011 |
| 60-6.4 | 1.000 | 1.000 | 1.10.4 | 1.090 | 1.012 | 1.031 |
| 65-69) | 1.000 | 1.000 | 1.146 | 1.100 | 1.082 | 1.072 |
| $70+$ | 1.000 | 1.000 | 1.114 | 1.000 | . 979 | 1.000 |
| Totals | 1.029 | 1.026 | 1.070 | 1.051 | 1.039 | 1.025 |

Note: The preliminary correction factors for censuses are calculated by se and age group by dividing the corrected populations of Table B. 1 by the corresponding reported populations given in Tables 1 and 2.

Table B.3. Coale-Demeny model North life tables for determining preliminary correction factors for registered deaths for males: Thailand, 1960-65, 1965-70, 1970-75, and 1975-80

| Agex | 1960-65 |  | 1965-70 |  | 1970-75 |  | 1975-30 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $l_{x}$ | $=1.3$ | 13 | $=L_{\text {a }}$ | $l_{x}$ | ${ }_{5} L_{x}$ | $l_{x}$ | ${ }_{5} L_{x}$ |
| 0 | 100,000 | 451.43 .3 | 100,000 | 453.753 | 100,000 | 456,072 | 100,000 | 458.392 |
| 5 | 87, 3.41 | 432.926 | 88,170 | +36,320 | 88,799 | 439,715 | 89.428 | 443,109 |
| 10 | 8ร. 630 | 425,476 | 86,358 | 429.240 | 87.087 | 433,005 | 87,815 | 436,769 |
| 15 | 84.301 | 419.212 | $8=.338$ | 423,207 | 86,115 | 427,201 | 86,892 | 431,197 |
| 20 | 83,124 | 410, 3 \% | 83.445 | 414,511 | 84,766 | 419,065 | 85.587 | 423,319 |
| 25 | S1,094 | 400.3 -4 | 81,979 | 404.908 | 82,860 | 409,461 | 83,740 | 414.015 |
| 30 | 79,043 | 384.843 | 79.984 | 394,698 | 80,925 | 399,553 | 81,866 | 404,408 |
| 35 | 7n, 84\% | 778.588 | 77.890 | 383,754 | 78,897 | 388.921 | 79,898 | 394.087 |
| 40 | 74.541 | .065,789 | 75.607 | 371,291 | 76,672 | 376.793 | 77,738 | 382,295 |
| 45 | 71.774 | 350,038 | 72.910 | 356,490 | 74,045 | 362,341 | 75,181 | 368,193 |
| 50 | nis.aso | 331.681 | 69.686 | 337.847 | 70.891 | 344.013 | 72,097 | 350,179 |
| 55 | 64.192 | 309805 | 65.453 | 314.230 | 66.714 | 320,672 | 67.975 | 327.106 |
| (6) | 58.401 | 277.125 | (r) . 2.43 | 283.736 | n1.955 | 290,347 | 62,867 | 296,958 |
| 65 | 51,919 | 230.904 | 53.251 | 243.480 | 54.583 | 250,051 | 55,915 | 256,623 |
| 70 | 42,844 | +12.342 | +4.141 | 429.776 | 45,438 | 447,210 | 46,735 | 464,644 |

[^2]Table B.4. Coale-Demeny model North life tables for determining preliminary correction factors for registered deaths for females: Thailand, 1960-65, 1965-70, 1970-75, and 1975-80

| Age $_{x}$ | 1960-65 |  | 1965-70 |  | 1970-75 |  | 1975-80 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $I_{x}$ | $L_{2}$ | 12 | ; $L_{2}$ | 12 | $L_{*}$ | $I_{x}$ | $s L_{x}$ |
| 0 | 100,000 | +59,370 | 100,000 | +62.864 | $100.0 \%$ | 466,359 | 100,000 | 470,935 |
| 5 | 89,462 | 43.154 | 90.425 | +48.401 | 91,388 | 453,648 | 42,6\% | 460,502 |
| 10 | 87,800 | 436,542 | 88.936 | +42.444 | 90,072 | +48,346 | 91,536 | +56,060 |
| 15 | 86,817 | 431.191 | 58,042 | 437.535 | 89.267 | +43,879 | 90.8 ¢ | +52,179 |
| 20 | $85.661)$ | 424,759 | 86.973 | 431.559 | 88,286 | 438.360 | 90, (\%): | +47,268 |
| 25 | 84,244 | 417.186 | 85.652 | +24,480 | 87,059 | 431,792 | 88,904 | 441,372 |
| 30 | 82,630 | 408.648 | 84, 144 | 416.517 | 85.678 | 424,386 | 87,645 | +34,726 |
| 35 | 80, 829 | 399.064 | 82.463 | +07,560 | 84.097 | 416.057 | 86,245 | 427,241 |
| 40 | 78,797 | 388,007 | 80.562 | 397.135 | 82,327 | 406,262 | 84,651 | 418,301 |
| 45 | 76,406 | 375.313 | 78,292 | 385,051 | 80.179 | 394.759 | 82,669 | 407,593 |
| 50 | 73,732 | 359, 85 | 75.729 | 370,217 | 77,726 | 380,449 | 80,368 | 394,013 |
| 55 | 70,263 | 340,249 | 72,359 | 351.013 | 74,455 | 361.776 | 77,237 | 376,091 |
| 60 | 65,837 | 313,593 | 68,047 | 324,895 | 70.256 | 336,198 | 73,199 | 351,290 |
| 65 | 59,600 | 276,099 | 61,912 | 287,767 | 6-4.223 | 299,435 | 67,317 | 315,098 |
| 70 | 50.839 | 526,512 | 53.195 | 562,554 | 55,551 | 598,297 | 58,722 | 647,334 |

[^3]Source: Same as Table B. 3 .

Table B.5. Preliminary correction factors for registered births and deaths: Thailand, 1960-65, 1965-70, 1970-75, and 1975-80

| Births deaths and age group | 1960-65 |  | 1965-70 |  | 1970-75 |  | 1975-80 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females | Males | Females | Males | Females |
| Births | 1.234 | 1.300 | 1.220 | 1.250 | 1.199 | 1.211 | 1.179 | 1.174 |
| Deaths |  |  |  |  |  |  |  |  |
| 0-4 | 1.799 | 1.808 | 2.194 | 2.045 | 2.534 | 2.325 | 3.088 | 2.624 |
| 5-9 | 1.604 | 1.543 | 1.724 | 1.537 | 2.001 | 1.637 | 2.177 | 1.596 |
| 10-14 | 1.460 | 1.609 | 1.522 | 1.612 | 1.587 | 1.556 | 1.751 | 1.578 |
| 15-19 | 1.682 | 1.594 | 1.601 | 1.568 | 1.503 | 1.478 | 1.424 | 1.393 |
| 20-24 | 1.746 | 1.187 | 1.795 | 1.388 | 1.685 | 1.549 | 1.455 | 1.582 |
| 25-29 | 1.773 | 1.286 | 1.667 | 1.318 | 1.602 | 1.645 | 1.503 | 1.827 |
| 30-34 | 1.485 | 1.191 | 1.449 | 1.252 | 1.396 | 1.313 | 1.385 | 1.603 |
| 35-39 | 1.304 | 1.117 | 1.209 | 1.158 | 1.228 | 1.250 | 1.128 | 1.301 |
| 4(1-44 | i. 162 | 1.095 | 1.175 | 1.180 | 1.108 | 1.224 | 1.045 | 1.234 |
| 45-49 | . 991 | 1.116 | 1.064 | 1.112 | 1.102 | 1.206 | . 982 | 1.128 |
| 50-54 | 1.042 | 1.183 | 1.012 | 1.124 | 1.093 | 1.173 | 1.085 | 1.228 |
| 55-59 | 1.044 | 1.260 | 1.041 | 1.213 | 1.040 | 1.178 | 1.079 | 1.200 |
| 60-64 | 1.113 | 1.364 | 1.055 | 1.235 | 1.092 | 1.264 | 1.019 | 1.203 |
| 65-69 | 1.287 | 1.596 | 1.251 | 1.480 | 1.219 | 1.459 | 1.211 | 1.445 |
| $70+$ | 1.45 | 1.671 | 1.485 | 1.562 | 1.417 | 1.456 | 1.352 | 1.514 |

[^4]
## APPENDIX C. FORMULAS FOR DERIVING THE LIFE TABLES

The following formulas were used to calculate the intercensal abridged life table for five-year age groups for the five-year period beginning 1 April of the vear $1960+5(i-1)(i=1,2,3,4)$ :

$$
\begin{aligned}
& { }_{5} L_{11}=\left(P_{i+1,11} / B\right)\left(5 l_{11}\right) \\
& { }_{5}^{\prime}-a+5={ }_{5} L_{a}\left(P_{i+1,1,+5} / P_{i, a}\right) \\
& T_{75}=\left({ }_{5} L_{71} P_{i+1,75+}\right) /\left(P_{i, 70+}-P_{i+1,75+}\right)
\end{aligned}
$$

where ${ }_{5} I_{a}=$ life table person-years lived between ages $a$ and $a+5 ; T_{75}=$ person-years lived after age 75; $P_{i, a}=$ corrected age $a$ to $a+5$ population on 1 April of the year $1960+5(i-1), i=1,2,3,4,5$ (see Tables 1,2 , and 11); $B=$ corrected births in the five-year period; and $l_{a}=$ life table survivors at exact age a. The second of these three formulas was applied iteratively.

All three formulas result from equating the survival ratios of the life table to those of the corrected data. In particular, the third is obtained by solving for $T_{75}$ in

$$
T_{75} /\left({ }_{5} L_{70}+T_{75}\right)=P_{i+1,75+} / P_{i, 71+}
$$

which equates the survival ratio of the life table with that of the corrected data for the open-ended age group.

Then

$$
l_{a+5}=\left(_{5} L_{a} / 5\right)\left(1-D_{a l}^{\star} / l_{, a}^{3}\right)
$$

where $D_{a}^{\star}$ denotes the corrected number of deaths during the five-year period occurring before exact age $a+5$ to those aged $a$ to $a+5$ at the beginning of the period.

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[^0]:    Sources: Same as Table 1.

[^1]:    Note: The entries here may be obtained by dividing the corrected births or deaths by the corresponding reported births or death, in Tables 5 and 6 .

[^2]:     life tables for $1904-67$ and 1424 - 0 , hearly interpolated, were used to determine preliminary correction factors for rewistered deaths for the age group 70 - for each five-year intercensal period.
    Source: Coale and Demens (1,

[^3]:    Note: Model North femaie life table levels are 17.0. 17.6. 18.2 and 190 for 1900-65, 1955-70, 1970-75, and 1975-80 respectively. SPC female life tables for $1964-67$ and $1974-76$, linearly interpolated, were used to determine the preliminary correction factors for registered deaths
    tor the age group $70+$ for each five-year intercensal period. tor the age group $70+$ for each five-yar intercensal period.

[^4]:    Niote: Preliminary correction hators for registered births and deaths are determined as described in Appendix A.

